

Topics:

- Kalman Filter
- Particle Filter
- Rejection Sampling
- Importance Sampling

Lectures 13, 14, beginning of 15

Friday, November 3, 2023

Discrete states review:

Filtering - Bayes' filter

Smoothing - sum-product / forward-backward on HMM

Most-likely state sequence - Viterbi

Continuous states (today):

Filtering: estimate $P(s_t | o_{1:t})$

- Kalman Filter when transition

and sensor models are linear with Gaussian noise.

for exact inference:

EKF, Unscented, etc. when

not linear - Gaussian as

approximate inference

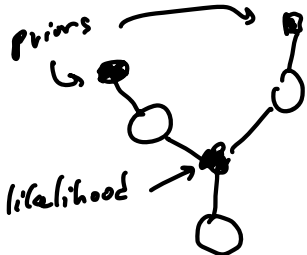
Kalman Filter

$$\underbrace{\alpha \cdot P(x)}_{\text{prior}} \cdot \underbrace{P(y|x)}_{\text{likelihood}} = \underbrace{P(x|y)}_{\text{posterior}}$$

x : belief

y : evidence

"conjugacy": When prior and likelihood are Gaussian, posterior is also Gaussian.



Bayes' Net or a factor graph.

On a Bayes' network \mathcal{Z} with linear - Gaussian distributions, the joint is multivariate Gaussian and conditioned on evidence, the cond. distr. is Gaussian.

Sampling from Continuous Distributions

• Rejection Sampling

$x \sim Q$

$z \sim \text{Uniform}[0, kQ(x)]$

if $z \leq P(x)$: accept. else reject.

\Rightarrow samples approximate P .

• Importance Sampling

$x \sim$ sample from Q .

weight sample by $\frac{P(x)}{Q(x)}$

\hookrightarrow statistics using weighted samples approximately statistics of P .

• Particle filter

Transition model $P(\underline{x}_t | \underline{x}_{t-1})$ same at each discrete time t and is linear - Gaussian
 Sensor model $P(y_t | \underline{x}_t)$ - - - - -
 Prior $P(\underline{x}_0)$

$\underline{x}, \underline{y}$, are vector-valued

Process Step

Probabilistic view.

$$P(\underline{x}_t | y_{0:t-1}) = \int P(\underline{x}_t | \underline{x}_{t-1}) \cdot P(\underline{x}_{t-1} | y_{0:t-1}) d\underline{x}_{t-1}$$

Since all Gaussian, can do linear algebra instead of calculus:

$$P(\underline{x}_t | y_{0:t-1}) = N(\underbrace{A\underline{x}_{t|t-1}}_{\hat{\underline{x}}_{t|t-1}}, \underbrace{A\underline{Q}_{t-1|t-1}A^T + W}_{\underline{Q}_{t|t-1}})$$

Transition Update

$$\underline{x}_t = A\underline{x}_{t-1} + \underline{w},$$

$\underline{w} \sim N(0, W)$
 independent of all r.v.s

Measurement Step

$$P(\underline{x}_t | y_{0:t}) \propto P(y_t | \underline{x}_t) \cdot P(\underline{x}_t | y_{0:t-1})$$

$$= N(\underbrace{\underline{x}_{t|t-1} + K_t [y_t - H\underline{x}_{t|t-1}]}_{\hat{\underline{x}}_{t|t}}, \underbrace{Q_{t|t-1} - K_t H Q_{t|t-1}}_{\underline{Q}_{t|t}})$$

Observation Update

$$y_t = H\underline{x}_t + \underline{v},$$

$\underline{v} \sim N(0, R)$
 independent of all r.v.s

where

K_t is 'Kalman gain', $K_t = Q_{t|t-1} H^T R^{-1}$ end slide 18 for Kalman gain in HW

Intuitively, it weights how much to consider y_t over $\underline{x}_{t|t-1}$.

Notice $Q_{t|t-1} \succ Q_{t-1|t-1}$ (transition) $Q_{t|t} \prec Q_{t|t-1}$ (observation).

Particle Filter

- Approximate inference
- The marginalize step as seen in Kalman Filter can quickly get cumbersome when dynamics aren't linear-Gaussian.

Ideas

- use samples as approximate representation of current state distribution.
- resample to focus on high-probability regions of state space.

Alg

- Init N particles from π .

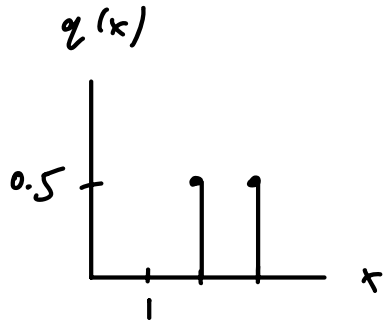
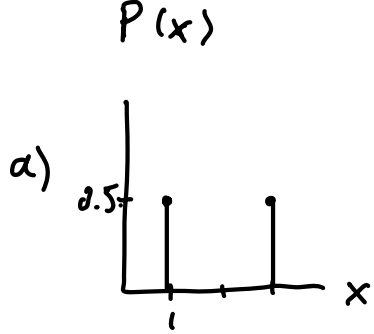
- Update:

sample $x_{t+1}^{[i]} \sim T(x_t^{[i]})$ for $i=1$ to N

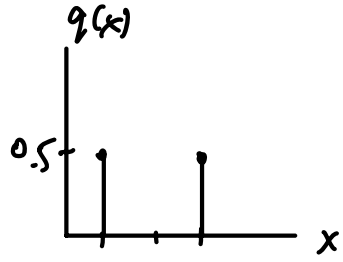
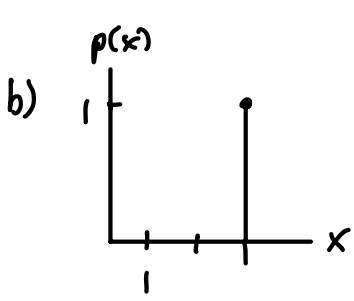
$w^{[i]} = P(y_{t+1} | x_{t+1}^{[i]})$ for $i=1$ to N

Generate x_{t+1} by sampling N elts from W .

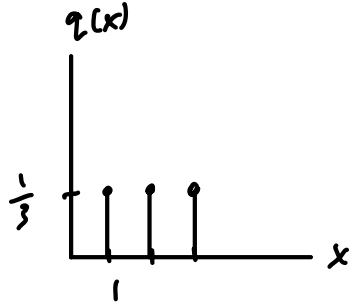
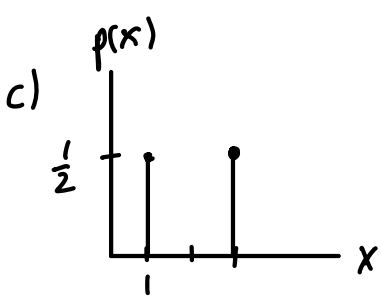
Answers



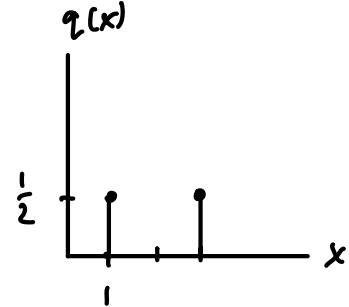
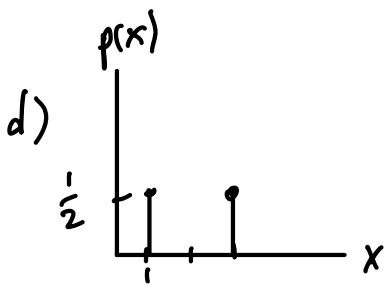
$D = \infty$



$D = 1$



$D = \log_2 3 - 1$



$D = 0$